

Practical statistical modeling of networks: Recent developments in exponential random graph (p^*) models

Co-chairs

Garry Robins, University of Melbourne

Martina Morris, University of Washington

Presenters

Pip Pattison, University of Melbourne

Mark S. Handcock, University of Washington

Garry Robins, University of Melbourne

Dave Hunter, Pennsylvania State University

Tom Snijders, University of Groningen

Steve Goodreau, University of Washington

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*1. Conceptual foundations
and some new model specifications*

Pip Pattison, University of Melbourne

or

**The left hand side of the equation learns to live
with the right hand side**

Three main points

We have learnt that:

1. We need slightly more subtle characterisations of dependencies among network ties than we initially thought

realisation-dependent not simply Markov

2. It is very helpful to simplify models by making (plausible) assumptions about relationships among related parameters

alternating k-stars/weighted degree distribution

alternating k-independent 2-paths/weighted (dyad-wise) shared partner distribution

alternating k-triangles/weighted (edge-wise) shared partner distribution

3. With these two insights, maximum likelihood estimation approaches work well

Siena (Snijders, 2002)

statnet (Handcock, Hunter, Butts, Goodreau & Morris, 2004)

What do we think we know about processes underlying network tie formation?

Exogenous effects:

Homophily, shared affiliations, spatial propinquity all matter (e.g., Marsden, 1988; Morris 1993; McPherson, Smith-Lovin & Cook, 2001)

More generally: network tie formation is often set within *foci* (Feld, 1981) or *settings* (White, 1995; Pattison & Robins, 2002)

Endogenous network effects:

Clustering: tie formation is often more likely when actors have network partners in common (e.g., Cartwright & Harary, 1956; Granovetter, 1973; and many others)

More generally: “a social tie ... is subject to, and known to be subject to, the hegemonic pressures of others engaged in the social construction of that network.” (White, 1998)

Interactions between exogenous and endogenous effects:

General social selection (Robins, Elliott & Pattison, 2001)

Modelling endogenous network processes

Guiding principles:

Network ties are the outcome of (unobserved) social processes that tend to be local and interactive

There are both regularities and irregularities in these local interactive processes

We hence construct statistical models in which:

local interactivity is permitted and assumptions about form of “local interactions” are explicit

regularities are represented by model parameters and estimated from data
consequences of local regularities for global network properties can be understood *and can also provide an exacting approach to model evaluation*

Local interactivity

We model *tie variables*: $\mathbf{X} = [X_{ij}]$ $X_{ij} = 1$ if i has a tie to j
0 otherwise

realisation of \mathbf{X} is denoted by $\mathbf{x} = [x_{ij}]$

We also incorporate node-level exogenous *attribute variables*: $\mathbf{Y} = [Y_i]$

Two modelling steps:

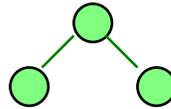
Local interactions: define two network tie variables to be *neighbours* if they are conditionally dependent, given the values of all other tie variables

But: what are appropriate *neighbour* assumptions?

Network topologies: which tie variables are neighbours?

Two tie variables are *neighbours* if:

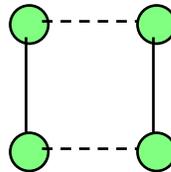
they share an actor



Markov model

(Frank & Strauss, 1986)

they share connections
with two existing ties
(completing a social circuit)



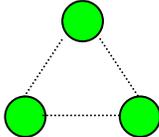
realisation-dependent model

(Pattison & Robins, 2002;
Snijders, Pattison, Robins &
Handcock, 2004)

There are other possibilities, but these two get us a long way

Models for interactive systems of variables (Besag, 1974)

A *neighbourhood* is a set of mutually neighbouring variables and corresponds to a potential *network configuration*:

e.g. $\{X_{12}, X_{13}, X_{23}\}$ corresponds to 

Hammersley-Clifford theorem (Besag, 1974):

A model for \mathbf{X} has a form determined by its neighbourhoods

This general approach leads to *exponential random graph* or *p* models*
(Frank & Strauss 1986; extended by Wasserman, Robins & Pattison)

Exponential random graph models

$$P(\mathbf{X} = \mathbf{x}) = (1/c) \exp\{\sum_Q \gamma_Q z_Q(\mathbf{x})\}$$

normalizing quantity

parameter

network statistic

the summation is over all neighbourhoods Q

$z_Q(\mathbf{x}) = \prod_{x_{ij} \in Q} x_{ij}$ signifies whether all ties in Q are observed in \mathbf{x}

$$c = \sum_{\mathbf{x}} \exp\{\sum_Q \gamma_Q z_Q(\mathbf{x})\}$$

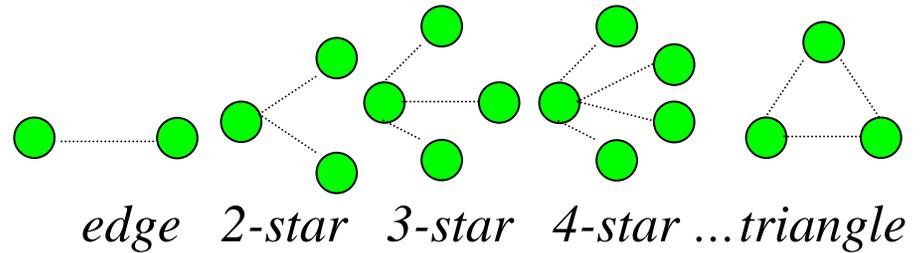
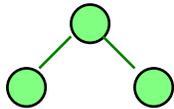
Neighbourhoods depend on proximity assumptions

Assumptions: two ties are *neighbours*:

Configurations for neighbourhoods

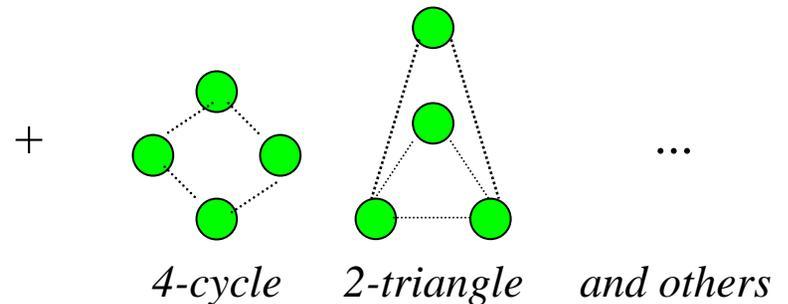
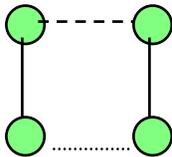
if they share an actor

Markov



if they complete a 4-cycle

realisation-dependent



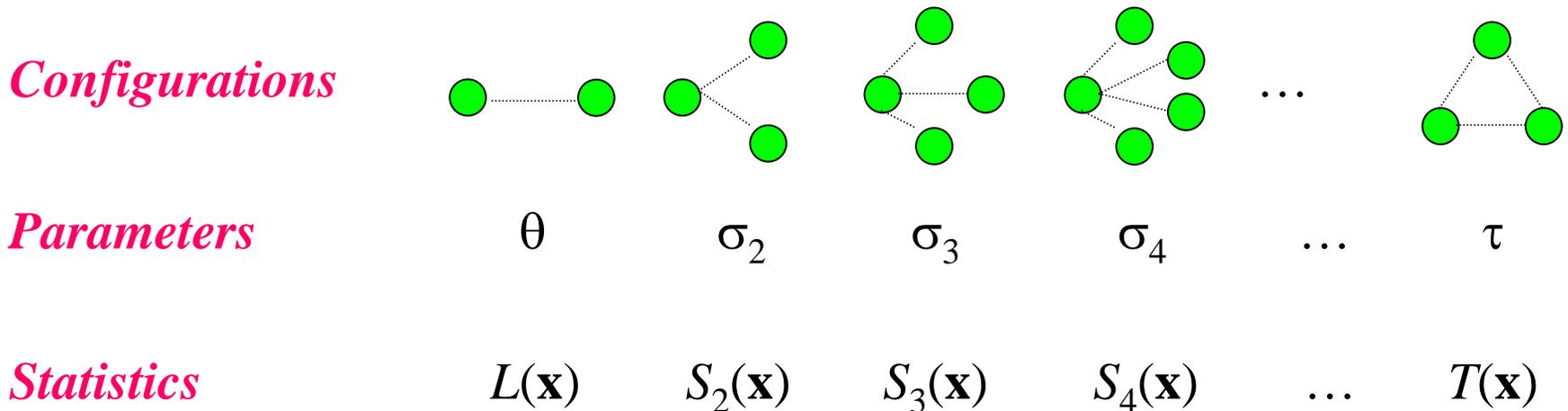
Homogenous models

If we assume that *isomorphic neighbourhoods have equal parameters*, then:

There is one parameter for *each class* of network configurations

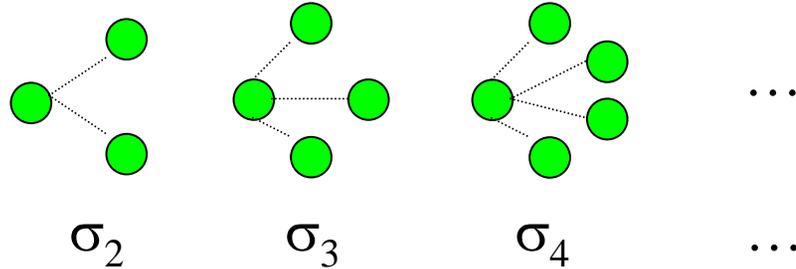
The corresponding statistic is the *number* of configurations in \mathbf{x}

E.g. for a Markov model:



Related model parameters

Star configurations



Parameters

If we assume that $\sigma_k = -\sigma_{k-1}/\lambda$, for $k > 1$ and $\lambda \geq 1$ a (fixed) constant *alternating k-star hypothesis*

Then we obtain a single *star* parameter (σ_2) with statistic:

$$S^{[\lambda]}(\mathbf{x}) = \sum_k (-1)^k S_k(\mathbf{x}) / \lambda^{k-2} \quad \textit{alternating k-star statistic}$$

Note that if $\lambda = 1$ and the edge parameter is included, the no. of isolated nodes is modelled separately

Geometrically decreasing degree distribution

An equivalent characterisation:

Consider statistics $d_k(\mathbf{x})$, where $d_k(\mathbf{x})$ is the number of nodes in \mathbf{x} of degree k (with corresponding parameters θ_k)

Assuming that

$$\theta_k = e^{-\alpha k} \quad \text{for } k = 1, 2, \dots, n-1$$

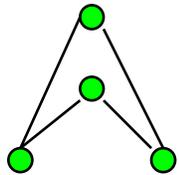
$$\lambda = e^\alpha / (e^\alpha - 1)$$

yields the statistic:

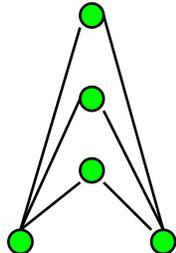
$$\sum e^{-\alpha k} d_k(\mathbf{x})$$

weighted degree distribution

Additional neighbourhoods for realisation-dependent model

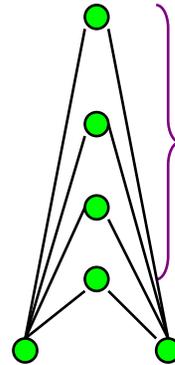


2-independent
2-path



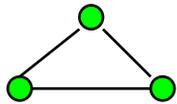
3-independent
2-path

...

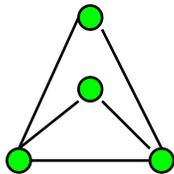


k -independent
2-path

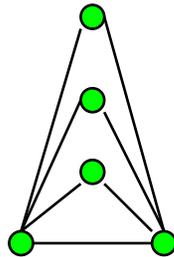
k nodes



triangle

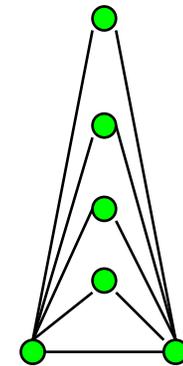


2-triangle



3-triangle

...



k -triangle

...

k nodes

Additional statistics

k-independent 2-paths

$U_k(\mathbf{x})$ = no of k -independent 2-paths in \mathbf{x} , with parameter v_k

Let $v_{k+1} = -v_k/\lambda$

Statistic for v_2 :

$$U^{[\lambda]}(\mathbf{x}) = \sum_k (-1)^k U_k(\mathbf{x})/\lambda^{k-2}$$

alternating independent 2-path statistic

k-triangles

$T_k(\mathbf{x})$ = no of k -triangles in \mathbf{x} , with parameter τ_k

Let $\tau_{k+1} = -\tau_k/\lambda$

Statistic for τ_1 :

$$T^{[\lambda]}(\mathbf{x}) = \sum_k (-1)^k T_k(\mathbf{x})/\lambda^{k-2}$$

alternating k-triangle statistic

Equivalently (Hunter & Handcock, 2004):

Statistic: No of dyads with exactly k shared partners

Aggregate statistic: *geometrically weighted dyad-wise shared partner distribution DSP*

Statistic: No of dyads linked by an edge and having exactly k shared partners

Aggregate statistic: *Geometrically weighted edge-wise shared partner distribution ESP*

Model specification: what have we learnt?

1. Relevant exogenous variables at node, tie and setting levels should be used
2. Realisation-dependent neighbourhoods appear to reflect social processes underlying network formation better than simple Markovian neighbourhoods
3. Hypotheses about relationships among the values of related parameters can provide practical and effective means of incorporating important higher-order effects without “death by parameter”
4. So we fit models with relevant *exogenous variables* as well as *edge, alternating 2-star, alternating k-independent 2-path and alternating k-triangle parameters* (**Siena**) or, equivalently, *edge, weighted degree distribution, DSP and ESP parameters* (**statnet**)

→ *Mark Handcock*